

ST790 — Fall 2022  
*Imprecise-Probabilistic Foundations of Statistics*

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Week 04a

- Possibility theory recap
- (Imprecise-)probability-to-possibility transform
- Extension principle

- A function  $\bar{\Pi} : 2^{\mathbb{X}} \rightarrow [0, 1]$  is a *possibility measure* if
  - $\bar{\Pi}(\emptyset) = 0$
  - $\bar{\Pi}(\mathbb{X}) = 1$
  - it's maxitive,<sup>1</sup> i.e.,  $\bar{\Pi}(\bigcup_{n=1}^{\infty} A_n) = \sup_n \bar{\Pi}(A_n)$
- There exists a function  $\pi : \mathbb{X} \rightarrow [0, 1]$ , called the *possibility contour*, such that  $\sup_{x \in \mathbb{X}} \pi(x) = 1$  and

$$\bar{\Pi}(A) = \sup_{x \in A} \pi(x), \quad A \subseteq \mathbb{X}$$

- The dual,  $\underline{\Pi}$ , is a *necessity measure* and satisfies

$$\underline{\Pi}(A) = 1 - \bar{\Pi}(A^c) = 1 - \sup_{x \in A^c} \pi(x), \quad A \subseteq \mathbb{X}$$

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<sup>1</sup>Update: For non-finite  $\mathbb{X}$ , I need to explicitly assume countable maxitivity in order to get the contour representation in 2nd bullet point

- $\bar{\Pi}$  determines a credal set  $\mathcal{C}(\bar{\Pi}) = \{P : P \leq \bar{\Pi}\}$ , containing all the probabilities *consistent* with it
- Characterization in terms of the sub-level sets of  $\pi$
- Consequences:
  - $\mathcal{C}(\bar{\Pi}) \neq \emptyset \implies$  so no-sure-loss
  - $\bar{\Pi}(\cdot) = \sup_{P \in \mathcal{C}(\bar{\Pi})} P(\cdot) \implies$  coherent<sup>2</sup>

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<sup>2</sup>*An aside:* There's a risk of incoherence, e.g., if we're not careful when "combining" distinct possibility measures about a common uncertain quantity, and we'll talk more about this later

## Proof.

To show: *for each  $A$ , there exists  $P \in \mathcal{C}(\bar{\Pi})$  such that  $P(A) = \bar{\Pi}(A)$ .*

Suppose  $\bar{\Pi}(A) < 1$ . Choose  $(x_n^\in) \subset A$  and  $(x_n^\notin) \subset A^c$  such that

$$\pi(x_n^\in) \rightarrow \bar{\Pi}(A) \quad \text{and} \quad \pi(x_n^\notin) \rightarrow \bar{\Pi}(A^c).$$

Define the sequences

$$\begin{aligned} g_n^\in &= \bar{\Pi}(\{x_1^\in, \dots, x_n^\in\}), & g_0 &= 0 \\ g_n^\notin &= \bar{\Pi}(A \cup \{x_1^\notin, \dots, x_n^\notin\}), & g_0^\notin &= \bar{\Pi}(A). \end{aligned}$$

Define a discrete probability  $P$ , with masses only on the above points, as

$$p(x_n^\in) = g_n^\in - g_{n-1}^\in \quad \text{and} \quad p(x_n^\notin) = g_n^\notin - g_{n-1}^\notin.$$

Then  $P$  is a probability, it's in  $\mathcal{C}(\bar{\Pi})$ , and  $P(A) = \bar{\Pi}(A)$ ..... □

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<sup>3</sup>From Hose's thesis, attributed to Fetz & Oberguggenberger (2004)

# Probability-to-possibility transform

- Suppose we have a probability  $P$  on  $\mathbb{X}$
- Goal: approximate<sup>4</sup>  $P$  by a possibility measure  $\bar{P} \geq P$
- A complement to the question we considered before:
  - instead of asking which  $P$  are compatible with  $\bar{P}$
  - we're asking which  $\mathcal{C}(\bar{P})$  contains a given  $P$ ?
- The *probability-to-possibility transform* says how to do this
- Of course, there's not a unique  $\bar{P} \geq P$ ...

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<sup>4</sup>The same approximation strategy applies when  $P$  is replaced by a general imprecise probability  $\bar{P}$ , and we'll consider this later

## Prob-to-poss transform, cont.

- Let  $f : \mathbb{X} \rightarrow \mathbb{R}$  be a (measurable) function
- For the given  $P$ , define a contour function

$$\pi_f(x) = P\{f(X) \leq f(x)\}, \quad x \in \mathbb{X}$$

- Define the corresponding possibility measure

$$\begin{aligned}\bar{\Pi}_f(A) &= \sup_{x \in A} \pi_f(x) \\ &= P\left\{f(X) \leq \sup_{x \in A} f(x)\right\}, \quad A \subseteq \mathbb{X}\end{aligned}$$

### Theorem.

For any  $f$  as above,  $\bar{\Pi}_f \geq P$  or, in other words,  $\mathcal{C}(\bar{\Pi}_f) \ni P$

## Proof.

At least two different proofs:

- For any  $A \subseteq \mathbb{X}$ , it's clear that

$$E := \{y : f(y) \leq \sup_A f\} \supseteq A$$

So,  $P(A) \leq P(E) = \bar{P}(A)$ , hence  $P \leq \bar{P}_f$

- Define the sub-level set of  $\pi_f$ :

$$S_\alpha^c = \{x : P[f(X) \leq f(x)] \leq \alpha\}$$

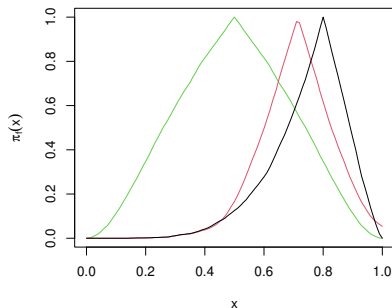
The right-hand side consists of those  $x$  such that  $f(x)$  is no larger than the lower- $\alpha$  quantile of  $f(X)$ . That event has probability  $\leq \alpha$ , so  $P(S_\alpha^c) \leq \alpha$  and  $P \in \mathcal{C}(\bar{P}_f)$





# Example

- $P = \text{Beta}(a, b)$  with  $a = 5$  and  $b = 2$
- Induced  $\pi$  based on three different  $f$ 
  - $f(x) = x(1-x)$
  - $f(x) = -|x - \frac{a}{a+b}|$
  - $f(x) = \text{dbeta}(x, a, b)$



## Example, cont.

- FYI, I'm doing this via Monte Carlo
- R code I used is below

```
63 # Probability-to-possibility transform
64
65 a <- 5; b <- 2
66 m <- a / (a + b)
67 N <- 5000
68 X <- rbeta(N, a, b)
69 f1 <- function(x) x * (1 - x)
70 f2 <- function(x) -abs(x - m)
71 f3 <- function(x) dbeta(x, a, b)
72 p1 <- function(x) sapply(x, function(y) mean(f1(X) <= f1(y)))
73 p2 <- function(x) sapply(x, function(y) mean(f2(X) <= f2(y)))
74 p3 <- function(x) sapply(x, function(y) mean(f3(X) <= f3(y)))
75 curve(p1, xlim=c(0,1), ylab=expression(pi[f](x)), col=3)
76 curve(p2, add=TRUE, col=2)
77 curve(p3, add=TRUE)
```

- Natural question: *which  $f$  is best?*
- The so-called *specificity principle* suggests that a “best”  $f^*$ , if it exists, would be such that

$$\pi_{f^*}(x) \leq \pi_f(x) \quad \text{for all } x \text{ and all } f$$

- Too few constraints to solve this...
- Suppose we have a *subjective plausibility order*
  - determined by a function  $r$ ,
  - i.e.,  $x$  is no less plausible than  $y$  iff  $r(x) \leq r(y)$
- Then the *most specific* approximation of  $P$  by possibility measure (relative to the order  $r$ ) is with  $f^* = r$ .

- A bit frustrating: we're left with the choice of ordering  $r$
- Where does this come from?
- Good default approach, when  $P$  has density  $p$ :
  - Recall: probable  $\implies$  possible
  - Makes sense for  $\pi_f(x) > \pi_f(y)$  iff  $p(x) > p(y)$
  - So, take  $f^* = r = p$
- This is what's shown in the black line in above plot
- Reminiscent of some ideas in statistics...

- $X \sim N_q(\theta, I)$
- Default posterior:  $\theta \sim N(x, I)$
- Possibilistic approximation of the posterior

$$\pi_f(\vartheta) = P_{\theta|x}\{f(\theta) \leq f(\vartheta)\}$$

- Above strategy suggests taking  $f^* = p_x$ , post density, so

$$\begin{aligned}\pi_{f^*}(\vartheta) &= P_{\theta|x}\{p_x(\theta) \leq p_x(\vartheta)\} \\ &= \dots \\ &= P\{\text{ChiSq}(q) \geq \|\vartheta - x\|^2\} \\ &= \text{p-value of the LR test of } H_0 : \theta = \vartheta\end{aligned}$$

- This is a simple but important idea
- Suppose  $\bar{\Pi}_X$  is a possibility measure on  $\mathbb{X}$ , contour  $\pi_X$
- This quantifies uncertainty about an “uncertain variable”  $X$
- Suppose  $X$  is related to  $Y$  in  $\mathbb{Y}$  via  $c(X, Y) = 0$
- How to quantify uncertainty about  $Y$ ?
- *Extension principle* says

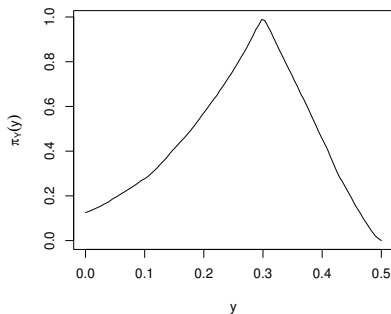
$$\bar{\Pi}_Y(B) = \sup_{y \in B} \pi_Y(y), \quad B \subseteq \mathbb{Y}$$

where

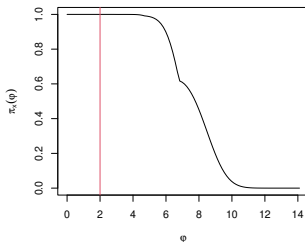
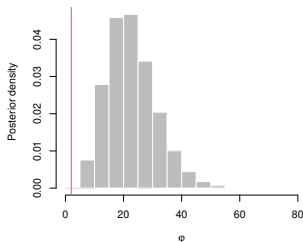
$$\pi_Y(y) = \sup_{x:c(x,y)=0} \pi_X(x), \quad y \in \mathbb{Y}$$

## Example, cont.

- $X$  is the uncertain quantity,  $\pi_X$  from before
- e.g., take  $Y = |X - \frac{1}{2}|$
- To evaluate  $\pi_Y(y)$  for a given  $y$ 
  - there's a pair  $x = \frac{1}{2} \pm y$
  - maximize  $\pi_X(x)$  over these two values



- Stein's example:  $X \sim N_q(\theta, I)$ , interest in  $\phi = \|\theta\|^2$
- Two solutions:
  - marginal Bayes (left), non-central chi-square
  - marginal possibility (right), via extension principle
- $q = 10$ , very naive computation<sup>5</sup> of  $\pi_x(\varphi)$
- Bayes misses true  $\phi$ , other doesn't



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<sup>5</sup>Semi-reliable:  $\pi_x(\phi) \approx 0.99$



- Interpretation of a possibility measure?
- At least two options:
  - interpret  $\bar{\Pi}$  directly
  - interpret  $\bar{\Pi}$  indirectly via  $\mathcal{C}(\bar{\Pi})$
- Shackle had in mind the former
- Latter might be more natural, e.g.,
  - if  $\bar{\Pi}(A)$  is small
  - then  $P(A)$  is small too for all compatible P's
- I don't know which is better/easier/etc.

- Belief functions
- Relations to random sets and possibility measures
- Properties
- ...