

ST790 — Fall 2022

Imprecise-Probabilistic Foundations of Statistics

Ryan Martin

North Carolina State University

www4.stat.ncsu.edu/~rmartin

Week 09b

- Towards a valid & efficient IM
 - Possibility measures and *strong validity*
 - Imprecise-probability-to-possibility transform
 - ...
-
- Presentation will be based on ideas in a recent paper¹ and some new stuff I'm working on now

¹<http://arxiv.org/abs/2203.06703>

■ Recall:

- uncertain $(Y, \Theta) \in \mathbb{Y} \times \mathbb{T}$, imprecise prob $\bar{P}_{Y, \Theta}$
- data $Y = y$ is observed, goal is inference on Θ
- an *inferential model* (IM) is a mapping $(y, \dots) \mapsto (\underline{\Pi}_y, \bar{\Pi}_y)$, output is an imprecise prob on \mathbb{T}
- an IM is *valid* (wrt $\bar{P}_{Y, \Theta}$) if

$$\bar{P}_{Y, \Theta} \{ \bar{\Pi}_Y(A) \leq \alpha, \Theta \in A \} \leq \alpha, \quad (\alpha, A) \in [0, 1] \times 2^{\mathbb{T}}$$

It is unacceptable if a procedure... of representing uncertain knowledge would, if used repeatedly, give systematically misleading conclusions (Reid & Cox 2015)

- Dempster's IM isn't valid
- Generalized Bayes is valid but inefficient
- What else can we do?
- ...
- Options:
 - modify Dempster's approach to force validity
 - leverage connections between possibility theory & validity

- Recall, Dempster's IM can incur sure-loss due to *contraction*, validity fails for the same reason
- Suggests Dempster's random sets, $\mathcal{T}_y(U)$, are "too small"
- How to enlarge them?
- Idea:² replace $U \sim P_U$ with a random set \mathcal{U} ,

$$\mathcal{T}_y(\mathcal{U}) = \bigcup_{u \in \mathcal{U}} \underbrace{\{\theta \in \mathbb{T} : y = a(\theta, u)\}}_{\mathcal{T}_y(u)}$$

- Union implies " $\mathcal{T}_y(\mathcal{U}) \supseteq \mathcal{T}_y(U)$ " in a stochastic sense
- Need some conditions on \mathcal{U} ...

²M. and Liu (2013, arXiv:1206.4091) and the book

Theorem (M. and Liu 2013).

Associate $Y = a(\Theta, U)$, with $U \sim P_U$; vacuous prior case.

Define $f(u) = P_{\mathcal{U}}(\mathcal{U} \ni u)$, a feature of $\mathcal{U} \sim P_{\mathcal{U}}$. If

- $f(U) \geq_{\text{st}} \text{Unif}(0, 1)$ when $U \sim P_U$,
- $\mathcal{T}_y(\mathcal{U})$ is non-empty for almost all y ,

then the IM with $\bar{P}_y(A) = P_{\mathcal{U}}\{\mathcal{T}_y(\mathcal{U}) \cap A \neq \emptyset\}$ is valid.

- Condition on \mathcal{U} is relatively mild, in particular, doesn't depend on special structure in the model
- Some recent³ work replaced \mathcal{U} with a possibility measure that dominates P_U

³<https://researchers.one/articles/20.08.00004>

Example

- $Y \sim \text{Bin}(n, \theta)$, vacuous prior
- $F_{n,\theta}(Y - 1) \leq 1 - U < F_{n,\theta}(U)$, with $U \sim P_U = \text{Unif}(0, 1)$
- How to choose \mathcal{U} to satisfy above condition?
 - set⁴ $h(u) = |u - \frac{1}{2}|$
 - and define $\mathcal{U} = \{u : h(u) \leq h(U)\}$
- Then⁵ the y -dependent random set for Θ is

$$\begin{aligned}\mathcal{T}_y(\mathcal{U}) &= \bigcup_{u \in \mathcal{U}} \{\theta : F_{n,\theta}(y - 1) \leq 1 - u < F_{n,\theta}(y)\} \\ &= \bigcup_{u \in \mathcal{U}} \{\theta : 1 - G_{y,n-y+1}(\theta) \leq 1 - u < 1 - G_{y+1,n-y}(\theta)\} \\ &= (\underline{\vartheta}_y(U), \bar{\vartheta}_y(U)]\end{aligned}$$

⁴Same result holds with virtually any other h

⁵Let $G_{a,b}$ denote the Beta(a, b) CDF

- Where

$$\begin{aligned}\underline{\vartheta}_y(U) &= G_{y+1, n-y}^{-1}\left(\frac{1}{2} - |U - \frac{1}{2}|\right) \\ \bar{\vartheta}_y(U) &= G_{y, n-y+1}^{-1}\left(\frac{1}{2} + |U - \frac{1}{2}|\right)\end{aligned}$$

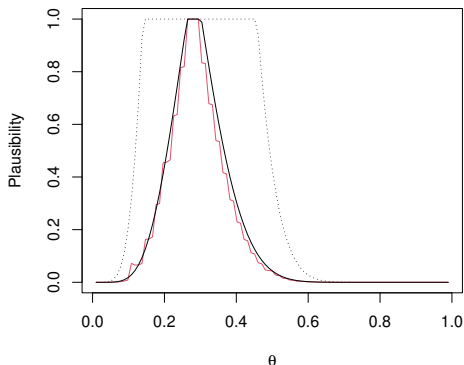
- IM's contour function

$$\begin{aligned}\pi_y(\theta) &= P_U\{\mathcal{T}_y(U) \ni \theta\} \\ &= 1 - P_U\{\underline{\vartheta}_y(U) > \theta\} - P_U\{\bar{\vartheta}_y(U) < \theta\} \\ &= \dots \\ &= 1 - \max\{0, 2G_{y+1, n-1}(\theta) - 1\} - \max\{0, 1 - 2G_{y, n-y+1}(\theta)\}\end{aligned}$$

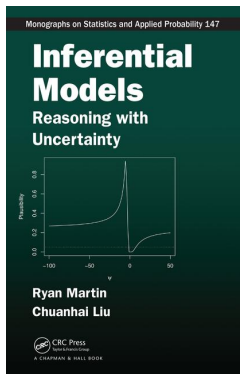
- Plot on the next page...

Example, cont.

- Data: $Y = 7, n = 25$
- Compare Walley (dotted), IM above, and **new IM**
- IM above is valid, more efficient than Walley, but slightly less efficient than the new IM



- Above result was the basis for the theory developed in the book
- I'm biased, but I think this is nice and interesting work
- Not perfect though:
 - requires " $Y = a(\Theta, U)$ "
 - requires user to specify \mathcal{U}
 - too complicated
 - doesn't handle partial priors
- Need to modernize the idea



- Validity aims at making probabilistic reasoning reliable
- In the vacuous-prior settings, this condition was also sufficient for constructing IM-based confidence regions⁶
- With genuine partial priors, validity isn't strong enough

Definition.

Let $(y, \dots) \mapsto (\underline{\Pi}_y, \overline{\Pi}_y)$ be an IM, and define the contour $\pi_y(\theta) = \overline{\Pi}_y(\{\theta\})$. This IM is *strongly valid* (wrt $\overline{P}_{Y,\Theta}$) if

$$\overline{P}_{Y,\Theta}\{\pi_Y(\Theta) \leq \alpha\} \leq \alpha, \quad \text{all } \alpha \in [0, 1]$$

⁶e.g., Theorem 1 in <https://researchers.one/articles/21.01.00002>

- Similar to Walley (2002)'s *fundamental frequentist principle*
- Makes the construction of IM-based confidence regions easy

$$C_\alpha(y) = \{\theta \in \mathbb{T} : \pi_y(\theta) > \alpha\}, \quad \alpha \in [0, 1]$$

- Implications (more details later):
 - some coherence-like properties
 - more-or-less forces the IM to be consonant
- Why stronger? Equivalent definition is

$$\bar{P}_{Y,\Theta}\{\underbrace{\bar{\Pi}_Y(A) \leq \alpha}_{\text{uniformity in } A}, \text{ for some } A \ni \Theta\} \leq \alpha, \quad \alpha \in [0, 1]$$

- How to achieve strong validity?

- Two key ingredients:⁷
 - outer consonant approximations
 - imprecise-probability-to-possibility transform
- Basic general facts:
 - Let \bar{P} be an upper prob for X and define
 - $\pi(x) = \bar{P}\{h(X) \leq h(x)\}$
 - $\bar{\Pi}(A) = \sup_{x \in A} \pi(x)$
 - then $\bar{\Pi}$ is an *outer consonant approximation* of \bar{P}
 - that is, $\bar{\Pi}$ is consonant and $\mathcal{C}(\bar{P}) \subseteq \mathcal{C}(\bar{\Pi})$
 - which implies $\bar{P}\{\pi(X) \leq \alpha\} \leq \alpha$
- Use these ideas to construct a strongly valid IM

⁷See the Week 04a lecture material

- The explanation of *how* to get from the above ideas/facts to my specific construction is lengthy⁸
- I'll spare you that explanation and jump to the conclusion
- Easiest for the precise-model-partial-prior case
- Notation:
 - $p_\theta(y)$ is the density/mass function of the (precise) model
 - $q(\theta) = \bar{P}_\Theta(\{\theta\})$ is the (imprecise) prior contour
 - the *relative likelihood* is

$$\eta(y, \theta) = \frac{p_\theta(y) q(\theta)}{\sup_{t \in \mathbb{T}} p_t(y) q(t)}, \quad \theta \in \mathbb{T}$$

- kind of like Bayes's formula...

⁸Very much influenced by the fundamental principles described in Section 2.3.2.1 in Hose's PhD thesis

- Recall that $\bar{P}_{Y,\Theta}$ is known
- Define a (consonant) IM with contour function⁹

$$\pi_y(\theta) = \bar{P}_{Y,\Theta}\{\eta(Y, \Theta) \leq \eta(y, \theta)\}, \quad \theta \in \mathbb{T}$$

- Consonance means the upper prob is defined via optimization

$$\bar{P}_y(A) = \sup_{\theta \in A} \pi_y(\theta), \quad A \subseteq 2^{\mathbb{T}}$$

- Can marginalize via extension principle...
- Outer consonant approximation \implies strong validity

Theorem.

The IM above is strongly valid wrt $\bar{P}_{Y,\Theta}$.

⁹just an imprecise-probability-to-possibility transform

- Generalized Bayes is valid but not strongly valid
- Other IMs are strongly valid, but tend to be inefficient¹⁰
- The only IM that achieves strong validity and efficiency?
- Why do I say “efficient”?
 - naive¹¹ construction in vacuous prior case gives contour

$$\pi_Y(\theta) = \sup_{\vartheta} P_{Y|\vartheta} \left\{ \frac{p_{\vartheta}(Y)}{\sup_t p_t(Y)} \leq \frac{p_{\theta}(y)}{\sup_t p_t(y)} \right\}$$

- relative likelihood is an approx pivot, so \sup_{ϑ} drops out
- RHS \approx the p-value for the optimal (?) likelihood ratio test

¹⁰e.g., Sec. 4.3 in <https://researchers.one/articles/21.05.00001>

¹¹“Naive” because there are principles that can/should be applied to improve efficiency — more on this later

- What's important is how this deals with partial prior info
- Recall:

$$\pi_y(\theta) = \bar{P}_{Y,\Theta}\{\eta(Y, \Theta) \leq \eta(y, \theta)\}, \quad \theta \in \mathbb{T}$$

with $\eta(y, \theta) = p_\theta(y) q(\theta) / \sup_t \{\dots\}$

- Then partial prior info enters in two ways:
 - through the definition of η (depends on q)
 - through the model $\bar{P}_{Y,\Theta}$
- This partial prior dependence creates opportunity for efficiency gain compared to vacuous-prior case
- Important: incorporating partial prior doesn't ruin validity!

- How to compute the $\bar{P}_{Y,\Theta}$ -probability that drives π_y ?
- If \bar{P}_Θ is consonant, then the Choquet integral simplifies:

$$\pi_y(\theta) = \int_0^1 \left[\sup_{\vartheta: q(\vartheta) > \beta} P_{Y|\vartheta} \{ \eta(Y, \vartheta) \leq \eta(y, \theta) \} \right] d\beta$$

- Suggests a simple/naive Monte Carlo strategy:

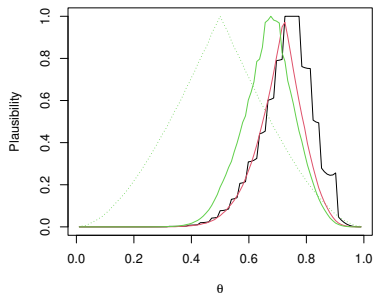
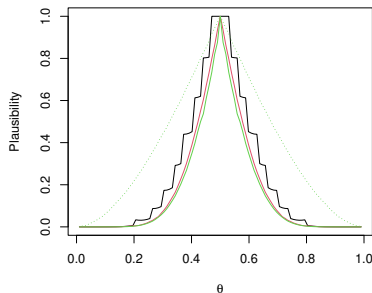
$$\pi_y(\theta) \approx \int_0^1 \max_{s: q(\vartheta_s) > \beta} \left[\frac{1}{M} \sum_{m=1}^M 1_{\{ \eta(Y_s^{(m)}, \vartheta_s) \leq \eta(y, \theta) \}} \right] d\beta$$

where

- $\vartheta_1, \dots, \vartheta_S$ is a fixed, dense grid
- $Y_s^{(1)}, \dots, Y_s^{(M)} \stackrel{\text{iid}}{\sim} P_{Y|\vartheta_s}$, for $s = 1, \dots, S$

Example

- $P_{Y|\theta} = \text{Bin}(n, \theta)$
- Vacuous, **complete**, and **partial** prior info¹²
- Partial prior info is *more* influential than complete



¹²There are dimension-reduction steps taken that I'm not explaining...

- New construction of a strongly valid IM
- Advantage is that it's based solely on the *posited model*
 - no choice of association
 - no choice of a random set (\mathcal{U})
- Partial prior is like regularization
 - allows for efficiency gains
 - without sacrificing validity
- Wanna see how this works in high-dim cases
- Need to scale up the computations

- Coherence-like properties
- Efficiency considerations
 - basic dimension-reduction
 - marginalization
- ...